

PARAMETRIC

ESTIMATORS

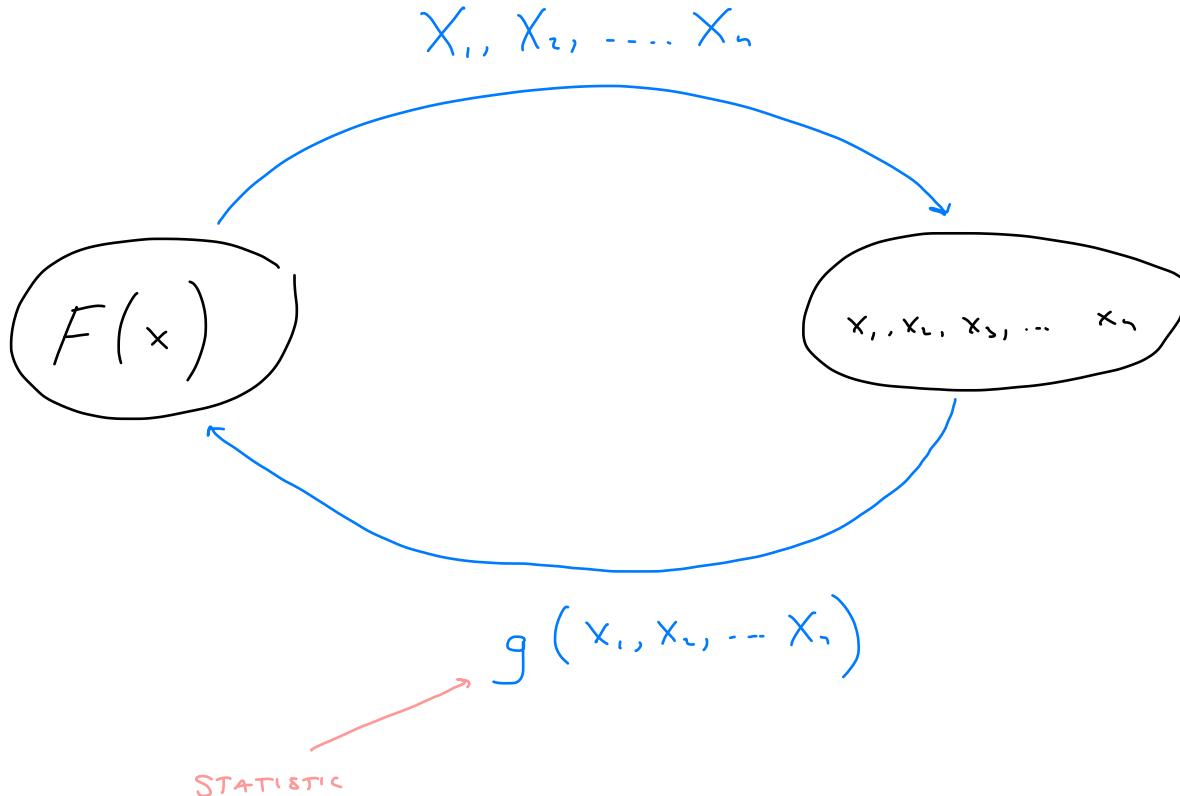
STAT 510

- PARAMETRIC MODELS
- CHECKING ASSUMPTIONS
- METHOD OF MOMENTS
- EXAMPLES
- LIKELIHOOD
- MAXIMUM LIKELIHOOD
 - EXAMPLES
 - EQUIVARIANT
 - OPTIMIZATION
- PARAMETRIC BOOTSTRAP

$$X_1, X_2, X_3, \dots, X_n \sim F$$

How do we "infer" F ?





↳ A FUNCTION OF DATA

↳ CAN BE A RANDOM VARIABLE

$$X_1, X_2, \dots, X_n \sim F$$

• How to estimate F ? STATISTICAL MODEL

• PARAMETRIC ESTIMATION



• NONPARAMETRIC ESTIMATION

$$X_1, X_2, \dots, X_n \sim F$$

PARAMETRIC
STATISTICAL
Model

$$\mathcal{F} = \left\{ f(x; \theta) : \theta \in \Theta \right\}$$

DENSITY PARAMETER THAT DETERMINES DENSITY PARAMETER SPACE

$$X_1, X_2, \dots, X_n \sim F$$

Assume $f \in \mathcal{F}$ where

$$\mathcal{F} = \left\{ f(x; p) = p^x (1-p)^{x-1}, \quad 0 < p < 1 \right\}$$

$$X_1, X_2, \dots, X_n \sim \text{Bernoulli}(p)$$

Problem reduced to estimating p

$$X_1, X_2, \dots, X_n \sim \text{Normal}(\mu, \sigma^2)$$

Assume $f \in \mathcal{F}$ where

$$\mathcal{F} = \left\{ f(x; \mu, \sigma^2) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}, \mu \in \mathbb{R}, \sigma > 0 \right\}$$

PARAMETER SPACE

IF PARAMETER
OF INTEREST

NUISANCE
PARAMETER

$$X_1, X_2, \dots, X_n \sim N(\mu, \sigma^2)$$

Point Estimation

USE $\hat{\theta}$ AS "BEST GUESS" FOR θ .

PARAMETER OF INTEREST

FIXED, UNKNOWN

GIVEN $X_1, X_2, \dots, X_n \sim F$

POINT ESTIMATOR $\longrightarrow \hat{\theta}_n = g(X_1, X_2, \dots, X_n)$

Random VARIABLE \nearrow

FUNCTION \uparrow

DATA \curvearrowright

SAMPLING DISTRIBUTION...

CHECKING

Assumptions

$$X_1, X_2, \dots, X_n \sim N(\mu, \sigma^2)$$



Assumption

- GRAPHICAL
- Formal Tests

METHOD OF MOMENTS

PARAMETER

$$\Theta = (\theta_1, \dots, \theta_k)$$

DISTRIBUTION OF THE PARAMETER IS A FUNCTION

jTH MOMENT

$$\alpha_j \equiv \alpha_j(\theta) = E_\theta [x^j] = \int x^j dF_\theta(x)$$

MOMENTS ARE FUNCTIONS OF THE PARAMETER

jTH SAMPLE MOMENT

$$\hat{\alpha}_j = \frac{1}{n} \sum_{i=1}^n x_i^j$$

METHOD OF MOMENTS

MoM

THE MoM ESTIMATOR $\hat{\theta}_n$ IS THE VALUE OF θ SUCH THAT

$$\alpha_1(\hat{\theta}_n) = \hat{\alpha}_1$$

SAMPLE SECOND MOMENT

$$\alpha_2(\hat{\theta}_n) = \hat{\alpha}_2$$

SECOND
MOMENTS

$$\alpha_k(\hat{\theta}_n) = \hat{\alpha}_k$$

K EQUATIONS, K UNKNOWNS \rightarrow SOLVE

Example

$X_1, X_2, \dots, X_n \sim \text{Geom}(p)$

$$\alpha_1 = \mathbb{E}[X] = 1/p$$

$$\hat{\alpha}_1 = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x}$$

$$1/\hat{p} = \bar{x} \Rightarrow \hat{p} = 1/\bar{x} \quad \checkmark$$



Mom ESTIMATOR

Example

$$X_1, \dots, X_n \sim N(\mu, \sigma^2)$$

$$\alpha_1 = \mathbb{E}[x] = \mu$$

$$\alpha_2 = \mathbb{E}[x^2] = \mathbb{V}[x] + (\mathbb{E}[x])^2 = \sigma^2 + \mu^2$$

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i \quad \Rightarrow \quad \hat{\mu} = \bar{x}_n$$

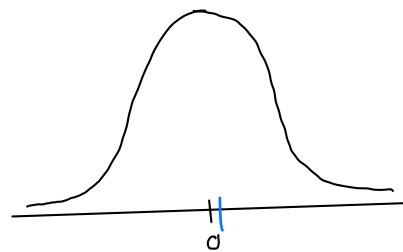
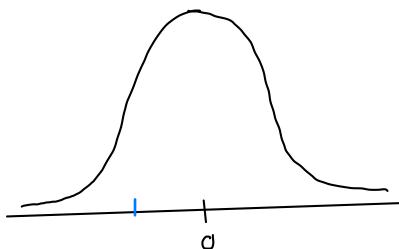
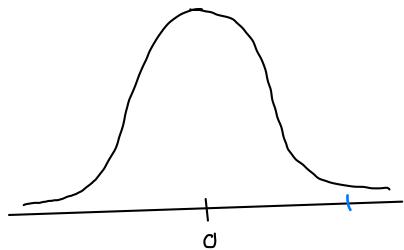
$$\hat{\sigma}^2 + \hat{\mu}^2 = \frac{1}{n} \sum_{i=1}^n x_i^2$$

$$\hat{\sigma}^2 = \frac{1}{n} \left(\sum x_i - \bar{x}_n \right)^2$$

LIKELIHOOD

Consider $X \sim N(0, 1)$

Known

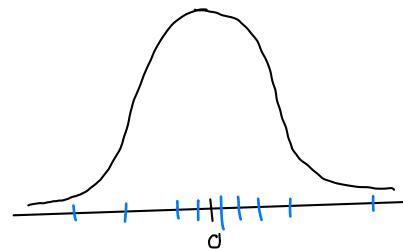
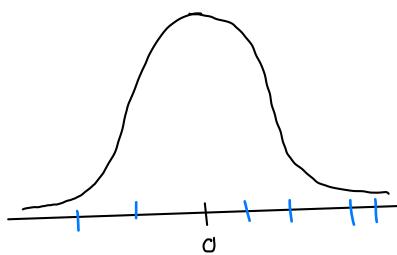
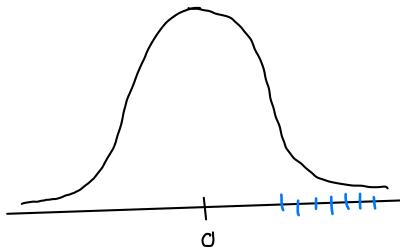


more likely to
be sampled



L I K E L I H O O D

Consider $X_1, \dots, X_n \sim N(\mu, \sigma^2)$
 μ known



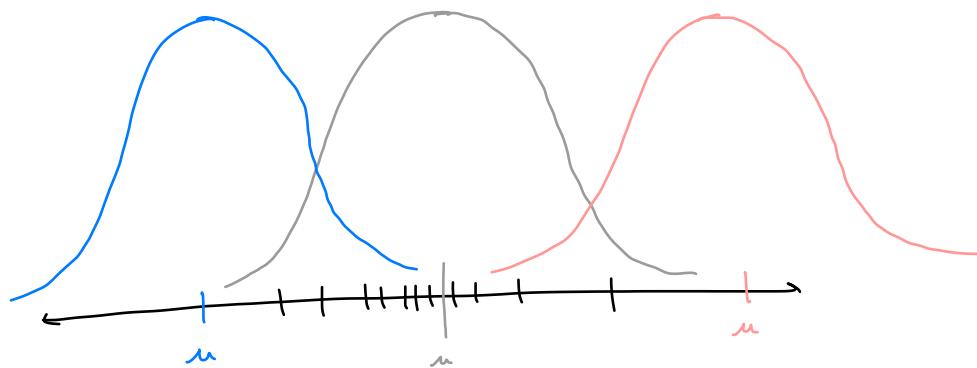
more likely to
be sampled



LIKELIHOOD

Assume $X_1, \dots, X_n \sim N(\mu, \sigma^2)$

UNKNOWN



Likelihood Function

X_1, \dots, X_n IID with PDF $f(x; \theta)$

$$\mathcal{L}_n(\theta) = f(x_1, x_2, \dots, x_n; \theta) = \prod_{i=1}^n f(x_i; \theta)$$

ASSUMES DATA OBSERVED

FUNCTION OF $\underline{\theta}$

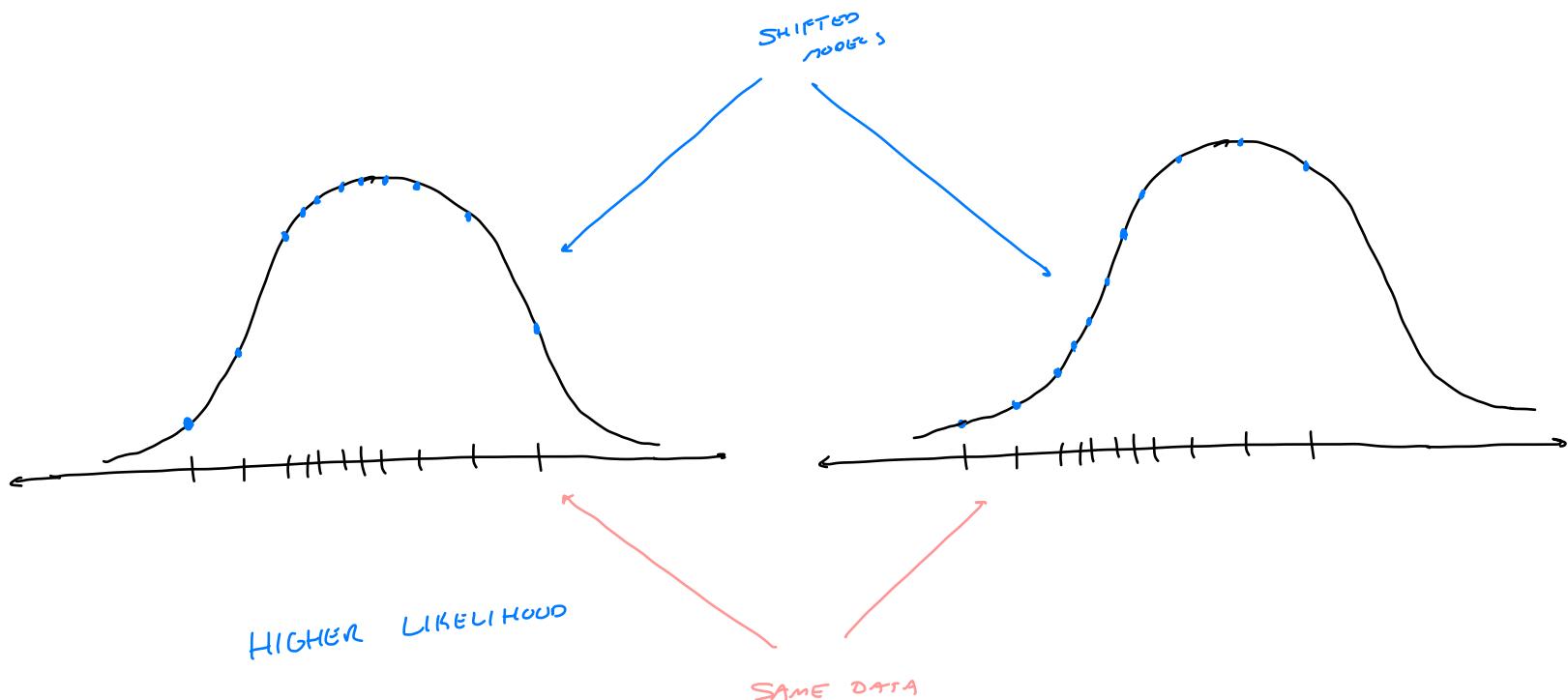
JOINT DENSITY IF θ KNOWN

DOES NOT NEED TO
INTEGRATE TO 1

$$\mathcal{L}_n(\theta) : \underline{\theta} \rightarrow [0, \infty)$$

Log-Likelihood

$$\ell_n(\theta) = \log \mathcal{L}_n(\theta)$$



Maximum Likelihood Estimation

MLE

$$\hat{\theta}_n = \underset{\theta}{\operatorname{argmax}} \quad L_n(\theta)$$

$$\hat{\theta}_n = \underset{\theta}{\operatorname{argmax}} \log L_n(\theta)$$

Example

$$X_1, \dots, X_n \sim \text{Geom}(p)$$

$$f(x; p) = (1-p)^{x-1} p \quad x=1, 2, \dots \quad p \in (0, 1)$$

$$\mathcal{L}(p) = \prod_{i=1}^n (1-p)^{x_i-1} p = p^{\sum x_i - n}$$

$$\log \mathcal{L}(p) = n \log p + (\sum x_i - n) \log(1-p)$$

$$\frac{1}{p} \log \mathcal{L}(p) = \frac{n}{p} - \frac{\sum x_i - n}{1-p} = 0$$

Some other checks

$$\Rightarrow \hat{p} = 1/\bar{x}$$

Example

$$X_1, \dots, X_n \sim N(\mu, \sigma^2)$$

$$\mathcal{L}(\mu, \sigma) = \prod_{i=1}^n \frac{1}{\sigma \sqrt{2\pi}} \exp \left[-\frac{1}{2} \left(\frac{x_i - \mu}{\sigma} \right)^2 \right]$$

two parameters

$$\begin{aligned} \frac{\partial}{\partial \mu} \log \mathcal{L}(\mu, \sigma) &= 0 \\ \frac{\partial}{\partial \sigma} \log \mathcal{L}(\mu, \sigma) &= 0 \end{aligned} \quad \left. \begin{array}{l} \\ \end{array} \right\} \text{Solve for } \mu, \sigma$$

Example

$$X_1, \dots, X_n \sim \text{UNIF}(0, \theta)$$

$$f(x; \theta) = \begin{cases} 1/\theta & 0 \leq x \leq \theta \\ 0 & \text{o.w.} \end{cases}$$

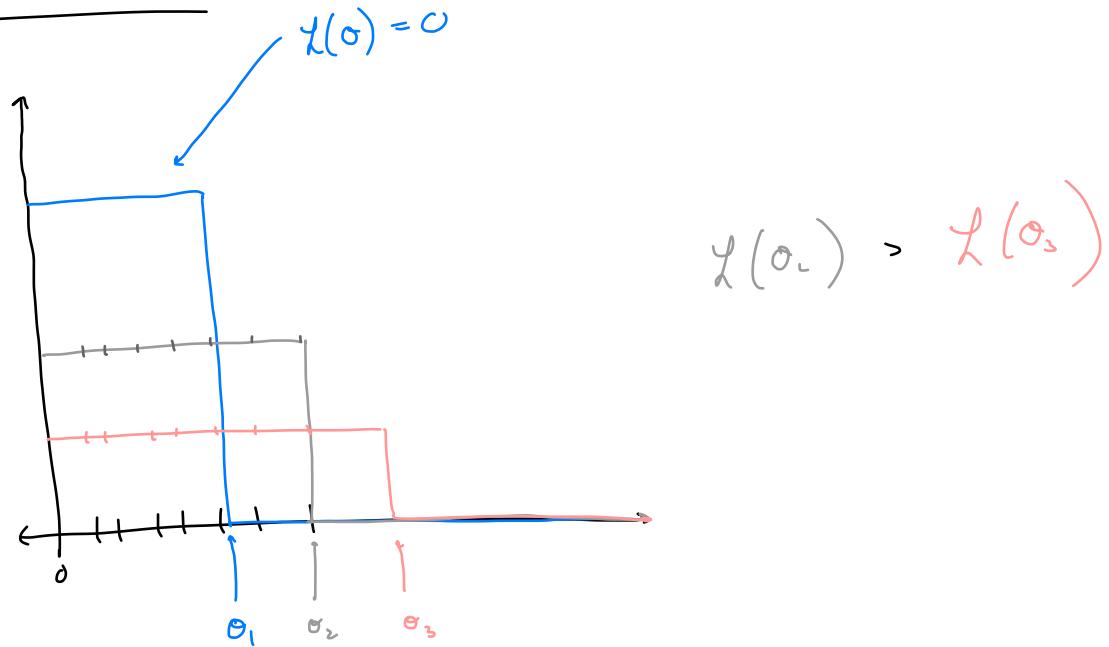
MAXIMUM IN DATA
WHICH IS KNOWN

$$\ell(\theta) = \begin{cases} (1/\theta)^n & \theta \geq X_{(n)} \\ 0 & \theta < X_{(n)} \end{cases}$$

AFTER $\theta = X_{(n)}$ $\ell(\theta)$ DECREASES

so $\hat{\theta} = X_{(n)}$

INTUITION



EQUIVARIANT

IF $\hat{\theta}_n$ IS MLE FOR θ

THEN $g(\hat{\theta}_n)$ IS MLE FOR $g(\theta)$

Example

$$X_1, \dots, X_n \sim N(\mu, 1)$$

$\hat{\mu} = \bar{X}$ is mle for μ

$e^{\hat{\mu}} = e^{\bar{X}}$ is mle for e^μ

OPTIMIZATION

$$\frac{\partial}{\partial \theta} \log \lambda(\theta) = 0$$

MIGHT NOT HAVE ANALYTIC SOLUTION

- NEWTON RAPHSON
- EM ALGORITHM

$$X_1, X_2, X_3, \dots, X_n \sim F$$

How do we "infer" F ?

LEARN



$$X_1, \dots, X_n \sim N(\mu, \sigma^2)$$

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2$$

LEARNED DISTRIBUTION

PARAMETRIC Bootstrap

NON-PARAMETRIC Bootstrap

- USE $\hat{F}_n(x)$ AS ESTIMATE OF $F(x)$
- SAMPLE FROM $\hat{F}_n(x)$

EMPIRICAL DISTRIBUTION

PARAMETRIC Bootstrap

- ASSUME PARAMETRIC FORM OF $F_0(x)$
- ESTIMATE PARAMETERS
- SAMPLE FROM LEARNED DISTRIBUTION, $\hat{F}_0(x)$

$$x_1^*, \dots, x_n^* \sim N(\hat{\mu}, \hat{\sigma}^2)$$